

# Online Appendix: Gibbs Sampler and VB

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## 1 The Model

$$y = Z\theta + \epsilon \tag{1}$$

where  $y$  is  $T \times 1$  vector of responses,  $Z$  is the  $T \times k$  matrix of standardized regressors,  $\epsilon$  is the  $T \times 1$  vector of *i.i.d* Normal errors with mean 0 and unknown variance  $\sigma^2$ .

## 2 The Priors

Following Bhattacharya, Pati, Pillai, and Dunson (2015) (BPPD), we elicit hierarchical DL prior as following:

$$\theta_j | \phi, \tau \sim DE(\phi_j \tau), \quad \phi_j \sim Dir(a, \dots, a) \tag{2}$$

Conditional prior for  $\theta_j$  is  $DE(\phi_j \tau)$  implies a zero mean Double Exponential or Lapalace distribution with the density  $f(\theta_j) = (2\phi_j \tau)^{-1} \exp(-\frac{|\theta_j|}{\phi_j \tau})$  for  $\theta_j \in \mathbb{R}$ .

Next, we set a Gamma priors for  $\tau$  and  $\sigma^{-2}$ :

$$\tau \sim G(ka, 1/2), \quad \sigma^{-2} \sim G(\nu, S). \tag{3}$$

The above hierarchical *DL* prior for  $\theta_j$  can be expressed as:

$$\theta_j \sim N(0, \psi_j \phi_j^2 \tau^2), \quad \psi_j \sim Exp(1/2) \tag{4}$$

Hence, the prior of  $\theta$  is  $N(\mathbf{0}, V)$  where  $V = \text{diag}(\psi_1\phi_1^2\tau^2, \dots, \psi_k\phi_k^2\tau^2)$ .

### 3 The Posteriors

Multiply the likelihood and the priors, we have

$$\begin{aligned}
& L(y|\theta, \sigma^2)p(\theta|\phi, \tau, \sigma^2)p(\tau)p(\phi)p(\sigma^2) \\
& \propto \frac{1}{(\sqrt{\sigma^2})^T} \exp\left[-\frac{1}{2\sigma^2}(y - Z\theta)'(y - Z\theta)\right] \times \prod_{j=1}^k \frac{1}{2\phi_j\tau} \exp\left[-\sum_{j=1}^k \frac{|\theta_j|}{\phi_j\tau}\right] \times \\
& \tau^{ka-1} \exp\left(-\frac{\tau}{2}\right) \times \prod_{j=1}^k \phi_j^{a-1} \times (\sigma^{-2})^{\nu-1} \exp(-S/\sigma^2)
\end{aligned} \tag{5}$$

or

$$\begin{aligned}
& L(y|\theta, \sigma^2)p(\theta|\phi, \tau, \psi, \sigma^2)p(\tau)p(\phi)p(\psi)p(\sigma^2) \\
& \propto \frac{1}{(\sqrt{\sigma^2})^T} \exp\left[-\frac{1}{2\sigma^2}(y - Z\theta)'(y - Z\theta)\right] \times \prod_{j=1}^k \frac{1}{\sqrt{\psi_j\phi_j^2\tau^2}} \exp\left[-\sum_{j=1}^k \frac{\theta_j^2}{2\psi_j\phi_j^2\tau^2}\right] \times \\
& \tau^{ka-1} \exp\left(-\frac{\tau}{2}\right) \times \prod_{j=1}^k \phi_j^{a-1} \times \prod_{j=1}^k \exp(-\psi_j/2) \times (\sigma^{-2})^{\nu-1} \exp(-S/\sigma^2)
\end{aligned} \tag{6}$$

In (6), the items involving  $\theta$  are  $\exp\left[-\frac{1}{2\sigma^2}(y - Z\theta)'(y - Z\theta)\right] \times \prod_{j=1}^k \frac{1}{\sqrt{\psi_j\phi_j^2\tau^2}} \exp\left[-\sum_{j=1}^k \frac{\theta_j^2}{2\psi_j\phi_j^2\tau^2}\right]$ . Thus, the conditional posterior of  $\theta$  is  $N(\bar{b}, \bar{V})$ , with  $\bar{V} = [\sigma^{-2}Z'Z + V^{-1}]^{-1}$  and  $\bar{b} = \sigma^{-2}\bar{V}Z'y$ .

In (5), the items involving  $\tau$  are  $\frac{1}{\tau^k} \exp\left[-\sum_{j=1}^k \frac{|\theta_j|}{\phi_j\tau}\right] \times \tau^{ka-1} \exp\left(-\frac{\tau}{2}\right)$ , which can be written as  $\tau^{ka-k-1} \exp\left[-\frac{1}{2}\left(\tau + \frac{1}{\tau}\left(\sum_{j=1}^k \frac{2|\theta_j|}{\phi_j}\right)\right)\right]$ . Thus the conditional posterior of  $\tau$  is  $giG(ka - k, 1, \sum_{j=1}^k \frac{2|\theta_j|}{\phi_j})$ .<sup>1</sup>

To find the conditional posteriors of  $\phi_j$ , we integrate  $\tau$  out following BPPD. Collecting the terms involving  $\phi_j$ , we have  $\prod_{j=1}^k (\phi_j^{a-1} \frac{1}{\phi_j}) \int_{\tau=0}^{\infty} \tau^{ka-k-1} \exp\left(-\frac{\tau}{2}\right) \exp\left[-\sum_{j=1}^k \frac{|\theta_j|}{\phi_j\tau}\right] d\tau$ . Thus, we can derive the conditional posterior of  $\phi_j$  as following: First, we have  $\xi_j \sim giG(a-1, 1, 2|\theta_j|)$ . Next let  $\Xi = \sum_{j=1}^k \xi_j$ . The conditional posterior of  $\phi_j$  can then be found to be  $\xi_j/\Xi$ .

In (6), the terms involving  $\psi_j$ s are  $\prod_{j=1}^k \psi_j^{-\frac{1}{2}} \exp\left[-\sum_{j=1}^k \frac{\theta_j^2}{2\psi_j\phi_j^2\tau^2}\right] \times \prod_{j=1}^k \exp(-\psi_j/2)$ . Thus

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<sup>1</sup> $y \sim giG(p, a, b)$  if  $f(y) \propto y^{p-1} \exp\left[-\frac{1}{2}(ay + b/y)\right]$ .

the conditional posterior of  $1/\psi_j$  is Inverse Gaussian with mean  $\sqrt{\frac{\phi_j^2 \tau^2}{\theta_j^2}}$  and scale parameter 1.

In (6), the terms involving  $\sigma^{-2}$  are

$$\frac{1}{(\sqrt{\sigma^2})^T} \exp\left[-\frac{1}{2\sigma^2}(y - Z\theta)'(y - Z\theta)\right] \times (\sigma^{-2})^{\nu-1} \exp(-S/\sigma^2)$$

Thus the conditional posterior of  $\sigma^{-2}$  is  $G(\frac{T}{2} + \nu, \frac{1}{2}(y - Z\theta)'(y - Z\theta) + S)$ .

## 4 Variational Bayes

The optimal VB  $q$  densities are as follows:

### 4.1 $q(\theta)$

$$q(\theta) \sim N(\bar{\theta}, \bar{V}), \tag{7}$$

where

$$\bar{V} = \left(\frac{T}{2} + \nu\right) Z'Z + V^{-1})^{-1}$$

$$\bar{\theta} = \left(\frac{T}{2} + \nu\right) \bar{V} Z' y$$

$$V^{-1} = \text{diag}(\overline{\psi_1^{-1} \phi_1^{-2} \tau^{-2}}, \dots, \overline{\psi_k^{-1} \phi_k^{-2} \tau^{-2}})$$

$$E[\log(|V^{-1}|)] = \sum_{j=1}^k E[\log(\psi_j) + 2 \log(\phi_j)] + 2k E[\log(\tau)],$$

where

$$E[\log(\psi_j) + 2 \log(\phi_j)] = \int_0^\infty q(\psi_j) \log(\psi_j) d\psi_j + 2 \int_0^\infty q(\phi_j) \log(\phi_j) d\phi_j$$

and

$$E[\log(\tau)] = \int_0^\infty q(\tau) \log(\tau) d\tau$$

#### 4.2 $q(\sigma^{-2})$

$$q(\sigma^{-2}) \sim G\left(\frac{T}{2} + \nu, \bar{S}\right), \quad (8)$$

where

$$\bar{S} = \frac{1}{2}[\|y - Z\bar{\theta}\|^2 + \text{tr}(Z'Z\bar{V})] + S$$

Hence

$$\bar{\sigma}^{-2} = \frac{\frac{T}{2} + \nu}{\bar{S}},$$

and

$$E[\log(\sigma^{-2})] = \psi\left(\nu + \frac{T}{2}\right) - \log(\bar{S})$$

where  $\psi(\bullet)$  is the Digamma function.

#### 4.3 $q(\tau)$

$$q(\tau) \sim giG\left[ka - k, 1, \sum_{j=1}^k 2(\bar{\theta}_j^2 + \bar{V}_{jj})^{1/2} \frac{1}{\phi_j}\right], \quad (9)$$

Let  $\chi = \sum_{j=1}^k 2(\bar{\theta}_j^2 + \bar{V}_{jj})^{1/2} \frac{1}{\phi_j}$ , we have

$$\bar{\tau} = \frac{\sqrt{\chi} K_{ka-k+1}(\sqrt{\chi})}{K_{ka-k}(\sqrt{\chi})}$$

and

$$\bar{\tau}^2 = \bar{\tau}^2 + \chi \left[ \frac{K_{ka-k+2}(\sqrt{\chi})}{K_{ka-k}(\sqrt{\chi})} - \left( \frac{K_{ka-k+1}(\sqrt{\chi})}{K_{ka-k}(\sqrt{\chi})} \right)^2 \right]$$

where  $K_*[\bullet]$  is the modified Bessel functions of the second kind.

#### 4.4 $q(\psi_j)$

$$q\left(\frac{1}{\psi_j}\right) \sim iG\left(\sqrt{\frac{\phi_j^2 \bar{\tau}^2}{\bar{\theta}_j^2 + \bar{V}^{jj}}}, 1\right), \quad (10)$$

Let  $\rho = \sqrt{\frac{\phi_j^2 \tau^2}{\theta_j^2 + \bar{V}_{jj}}}$ ,

$$\frac{\bar{1}}{\bar{\psi}_j} = \rho$$

and

$$\bar{\psi}_j = 1 + 1/\rho$$

Note that to calculate *ELBO*, we need to use the following optimal  $q$  density of  $\psi_j$ .<sup>2</sup>

$$q(\psi_j) = \left(\frac{\psi_j}{2\pi}\right)^{1/2} \exp\left\{-\frac{(\psi_j - \rho)^2}{2\rho^2\psi_j}\right\} \quad (11)$$

#### 4.5 $q(\phi_j)$

$$q(\xi_j) \sim giG(a - 1, 1, 2\sqrt{\theta_j^2 + (\bar{V}_{jj})^2}) \quad (12)$$

Let  $\varpi = 2\sqrt{\theta_j^2 + (\bar{V}_{jj})^2}$ , we have

$$\bar{\xi}_j = \frac{\sqrt{\varpi} K_a(\sqrt{\varpi})}{K_{a-1}(\sqrt{\varpi})},$$

and

$$var(\xi_j) = \varpi \left\{ \frac{K_{a+1}(\sqrt{\varpi})}{K_{a-1}(\sqrt{\varpi})} - \left[ \frac{K_a(\sqrt{\varpi})}{K_{a-1}(\sqrt{\varpi})} \right]^2 \right\}.$$

where  $var(\bullet)$  denotes the variance.

Scaling  $\xi$ , we have

$$\bar{\phi}_j = \frac{\bar{\xi}_j}{\sum^k \bar{\xi}_j},$$

and

$$\bar{\phi}_j^2 = \bar{\phi}_j + \frac{var(\xi_j)}{(\sum^k \bar{\xi}_j)^2}$$

Thus, the optimal  $q$  density of  $\phi_j$  takes the following form:

$$q(\phi_j) \sim giG\left[a - 1, \sum^k \bar{\xi}_j, (2\sqrt{\theta_j^2 + (\bar{V}_{jj})^2}) / (\sum^k \bar{\xi}_j)\right] \quad (13)$$

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<sup>2</sup>If  $x$  is distributed as  $f(x)$ , then  $y = 1/x$  is distributed as  $\frac{1}{y^2} f(\frac{1}{y})$ .

## 4.6 ELBO

The evidence lower bound (*ELBO*) is as follows:<sup>3</sup>

$$\begin{aligned}
ELBO &= E\{\log p(\mathbf{y}, \theta, \sigma^2, \phi, \tau, \psi)\} - E\{\log q(\theta, \sigma^2, \phi, \tau, \psi)\} \\
&= E\{\log p(\mathbf{y}|\theta, \sigma^2, \phi, \tau, \psi)\} + E\{\log p(\theta)\} + E\{\log p(\sigma^2)\} \\
&\quad + E\{\log p(\phi)\} + E\{\log p(\psi)\} + E\{\log p(\tau)\} \\
&\quad - E\{\log q(\theta)\} - E\{\log q(\sigma^2)\} - E\{\log q(\tau)\} - E\{\log q(\phi)\} - E\{\log q(\psi)\} + \dots
\end{aligned} \tag{14}$$

where

$$E\{\log p(\mathbf{y}|\theta, \sigma^2, \phi, \tau, \psi)\} = -\frac{T}{2} \log(2\pi) - \frac{T}{2} \left[ -\psi\left(\nu + \frac{T}{2}\right) + \log(\bar{S}) \right] - \frac{\bar{S} - S}{\left(\frac{\bar{S}}{\nu + \frac{T}{2}}\right)}, \tag{15}$$

$$E\{\log p(\theta)\} = -\frac{k}{2} \log(2\pi) - \frac{1}{2} E[\log |\mathbf{V}|] - \frac{1}{2} [\bar{\theta}' \mathbf{V}^{-1} \bar{\theta} + \text{tr}(\mathbf{V}^{-1} \bar{\mathbf{V}})], \tag{16}$$

$$E\{\log p(\sigma^{-2})\} = \nu \log S - \log \Gamma(\nu) - (\nu - 1) \left[ -\psi\left(\nu + \frac{T}{2}\right) + \log(\bar{S}) \right] - \left( \frac{S}{\frac{\bar{S}}{\nu + \frac{T}{2}}} \right), \tag{17}$$

$$E\{\log q(\theta)\} = -\left(\frac{k}{2} + \frac{k}{2}\right) \log(2\pi) + \frac{1}{2} \log |\bar{\mathbf{V}}_i|, \tag{18}$$

$$E\{\log q(\sigma^{-2})\} = -\left[\nu + \frac{T}{2} - \log(\bar{S}) + \log\left(\Gamma\left(\nu + \frac{T}{2}\right)\right) - \left(\nu + \frac{T}{2} - 1\right) \psi\left(\nu + \frac{T}{2}\right)\right]. \tag{19}$$

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<sup>3</sup>It is seen that the *ELBO* described below contains many constant terms. These terms need to be dropped from *ELBOs* in VB iterations to speed up the process.

Notice that

$$\begin{aligned}
& E\{\log p(\mathbf{y}|\theta, \sigma^2, \phi, \tau, \psi)\} + E\{\log p(\theta)\} + E\{\log p(\sigma^2)\} - E\{\log(q(\theta))\} - E\{\log q(\sigma^2)\} \\
&= \frac{k}{2} - \frac{T}{2} \log(2\pi) + \frac{1}{2} \log(|\bar{V}|) - \frac{1}{2} E[\log(|V|)] - \frac{1}{2} [\bar{\theta}' V^{-1} \bar{\theta} + \text{tr}(V^{-1} \bar{V})] + \nu \log(\underline{s}) - \log \Gamma(\nu) \\
&- (\nu + \frac{T}{2}) \log(\bar{s}) + \log(\nu + \frac{T}{2}) \\
&= \frac{1}{2} \log(|\bar{V}|) - \frac{1}{2} E[\log(|V|)] - \frac{1}{2} [\bar{\theta}' V^{-1} \bar{\theta} + \text{tr}(V^{-1} \bar{V})] - (\nu + \frac{T}{2}) \log(\bar{s}) + \text{Const.}
\end{aligned} \tag{20}$$

Above terms are the baseline *ELBO* discussed in GKP2018. The additional terms are as follows:

$$\begin{aligned}
E\{\log p(\tau)\} &= -\log \Gamma(ka) - ka \left[ \int_0^\infty (q(\tau) \log \tau) d\tau \right] - 0.5\bar{\tau} \\
&= -ka \left[ \int_0^\infty (q(\tau) \log \tau) d\tau \right] - 0.5\bar{\tau} + \text{Const.}
\end{aligned} \tag{21}$$

$$E\{\log p(\psi)\} = \sum_{j=1}^k \left( \log \frac{1}{2} - 0.5\bar{\psi}_j \right) = -\frac{1}{2} \sum_{j=1}^k \bar{\psi}_j + \text{Const.} \tag{22}$$

$$\begin{aligned}
E\{\log p(\phi)\} &= \log \Gamma(ka) - k \log \Gamma(a) - (a-1) \sum_{j=1}^k \left[ \int_0^\infty (q(\phi_j) \log \phi_j) d\phi_j \right] \\
&= -(a-1) \sum_{j=1}^k \left[ \int_0^\infty (q(\phi_j) \log \phi_j) d\phi_j \right]
\end{aligned} \tag{23}$$

$$E\{\log q(\tau)\} = \int_0^\infty q(\tau) \log q(\tau) d\tau. \tag{24}$$

$$E\{\log q(\psi)\} = \sum_{j=1}^k \int_0^\infty q(\psi_j) \log q(\psi_j) d\psi_j \tag{25}$$

$$E\{\log q(\phi)\} = \sum_{j=1}^k \int_0^\infty q(\phi_j) \log q(\phi_j) d\phi_j \tag{26}$$

Terms in *ELBO* that do not have clear analytical forms can be numerically estimated.

## References

- [1] Bhattacharya, A., Pati, D., Pillai, N., and Dunson, D. (2015), Dirichlet-laplace priors for optimal shrinkage, *Journal of the American Statistical Association*, 110, 1479-1490.